# Stochastic Processes and Stochastic Calculus - 3 Markov Chains 

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## Overview

1 Discrete time processes

- An example
- The Markov property
- Discrete Markov chains

■ Invariant distributions
■ Irreducible and regular chains

- Ergodic theorems

2 Continuous time jump processes

- Continuous time Markov chains

■ Poisson process

## A "lake model" of employment flows

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■ flows between unemployment and employment
■ employment and unemployment rates at "equilibrium"

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Fix a unit of time (e.g. one month) and introduce two parameters
$\lambda:=$ probability that a worker loses his/her job in a month
$\phi:=$ probability that an unemployed one finds a job in a month
We call $\lambda, \phi \in[0,1]$ transition probabilities.

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We have

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\begin{array}{ll}
\mathbb{P}\left(X_{n+1}=E \mid X_{n}=E\right)=1-\lambda & \mathbb{P}\left(X_{n+1}=U \mid X_{n}=E\right)=\lambda \\
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NO. For example, we do not know the initial state $X_{0}$.

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But what about three or more? One year? At this stage we introduce the following assumption:

## Markov property

At any time, regardless of the information about the past months, the next-month state of employment depends uniquely on the present one, i.e.

$$
\begin{aligned}
& \mathbb{P}\left(X_{n+1}=E \mid X_{n}=U, X_{n-1}=\cdot, X_{n-2}=\cdot, \ldots, X_{n-k}=\cdot\right)=\mathbb{P}\left(X_{n+1}=E \mid X_{n}=U\right) \\
& \mathbb{P}\left(X_{n+1}=U \mid X_{n}=E, X_{n-1}=\cdot, X_{n-2}=\cdot, \ldots, X_{n-k}=\cdot\right)=\mathbb{P}\left(X_{n+1}=U \mid X_{n}=E\right) .
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- The model becomes very tractable analytically

■ It fits real data (?) - we can make predictions
■ Given only the data $\lambda$ and $\phi$ (from real world), this is the "fairest" model that one can think.
A more realistic assumption would be e.g.

$$
\mathbb{P}\left(X_{n+1}=E \mid X_{n}=U, X_{n-1}=U, X_{n-2}=U, \ldots, X_{n-k}=U\right) \leq \mathbb{P}\left(X_{n+1}=E \mid X_{n}=U\right)
$$

but how to make it quantitative?

## Examples



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$$
\begin{aligned}
\mathbb{P}\left(X_{2}=E, X_{1}=U \mid X_{0}=E\right) & =\mathbb{P}\left(X_{2}=E \mid X_{1}=U, X_{0}=E\right) \mathbb{P}\left(X_{1}=U \mid X_{0}=E\right) \\
& =\mathbb{P}\left(X_{2}=E \mid X_{1}=U\right) \mathbb{P}\left(X_{1}=U \mid X_{0}=E\right) \\
& =\phi \lambda
\end{aligned}
$$

$$
\mathbb{P}\left(X_{3}=E, X_{2}=U, X_{1}=U \mid X_{0}=E\right)=\phi(1-\phi) \lambda
$$

$$
\mathbb{P}\left(X_{3}=E \mid X_{0}=E\right)=?
$$

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Given a filtration $\left(\mathcal{F}_{t}\right)_{t \in T}$, a process $X_{t}: \Omega \rightarrow E$ is Markov if

$$
\text { for all } s \leq t \in T, A \subseteq E \quad \mathbb{P}\left(X_{t} \in A \mid \mathcal{F}_{s}\right)=\mathbb{P}\left(X_{t} \in A \mid X_{s}\right)
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■ Usually one has $\mathcal{F}_{t}=\sigma\left(X_{r}: r \leq t\right)$, the natural filtration of $X_{t}$.

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The transition probabilities are $|\mathcal{S}|^{2}$ numbers that we write matrix notation:

$$
\left(\begin{array}{ll}
1-\lambda & \lambda \\
\phi & 1-\phi
\end{array}\right) \Rightarrow Q=\left(p_{i j}\right)_{i, j \in \mathcal{S}} \quad p_{i j}=\mathbb{P}\left(X_{n+1}=j \mid X_{n}=i\right) \quad \text { for } i, j \in \mathcal{S}
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## Remark ( Q is a stochastic matrix)

We must have $p_{i j} \in[0,1]$ for $i, j \in \mathcal{S}$ and

$$
\sum_{j \in \mathcal{S}} p_{i j}=\sum_{j \in \mathcal{S}} \mathbb{P}\left(X_{n+1}=j \mid X_{n}=i\right)=\mathbb{P}\left(X_{n+1} \in \mathcal{S} \mid X_{n}=i\right)=1
$$

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we say that a process $X_{n}: \Omega \rightarrow \mathcal{S}(n \in \mathbb{N})$ is a markov chain if for all $n \in \mathbb{N}$,

$$
\begin{gathered}
\mathbb{P}\left(X_{n+1}=j \mid \mathcal{F}_{n}\right)=\mathbb{P}\left(X_{n+1}=j \mid X_{n}\right) \quad \forall j \in \mathcal{S} \quad \text { (Markov property) } \\
\mathbb{P}\left(X_{n+1}=j \mid X_{n}=i\right)=p_{i j} \quad \forall i, j \in \mathcal{S} \quad \text { (transition probability) }
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The two conditions above yield

$$
\mathbb{P}\left(X_{n+1}=j \mid X_{n}=i, X_{n-1}=\cdot, X_{n-2}=\cdot, \ldots, X_{n-k}=\cdot\right)=p_{i j}
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## Graphical representation

We associate
■ a node to each state $i \in \mathcal{S}$
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\mathcal{S}=(I, U, E, O) \\
Q=\left(\begin{array}{lll}
1-\eta & \eta & 0 \\
0 & \phi & 1-\phi \\
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## Theorem

Given a finite or countable $\mathcal{S}$, a stochastic matrix $Q=\left(p_{i j}\right)_{i, j \in \mathcal{S}}$ and a probability measure $\mu_{0}$ on $\mathcal{S}$.
$\exists)$ there exists a Markov chain $\left(X_{n}\right)_{n \in \mathbb{N}}$ (with respect to the natural filtration) with transition probability $Q$ and law of $X_{0}$ equal to $\mu_{0}$.
!) Such a chain $X$ is unique in law, i.e. given any Markov chain $Y$ with transition probability $Q$ and law of $X_{0}$ equal to $\mu_{0}$, one has

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The condition $X=Y$ in law means that for every $n \geq 1, A_{0}, \ldots, A_{n} \subseteq \mathcal{S}$

$$
\mathbb{P}\left(X_{0} \in A_{0}, X_{1} \in A_{1}, \ldots, X_{n} \in A_{n}\right)=\mathbb{P}\left(Y_{0} \in A_{0}, Y_{1} \in A_{1}, \ldots, Y_{n} \in A_{n}\right)
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Marginal laws

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The simplest case is that of 1-marginals, i.e. the law of $X_{n}$. For example

$$
P\left(X_{1}=j\right)=\sum_{i \in \mathcal{S}} P\left(X_{1}=j \mid X_{0}=i\right) P\left(X_{0}=i\right)=\sum_{i \in \mathcal{S}} p_{i j} \mu_{i}^{0}
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Recall the matrix-vector products for column vectors $v$ or row vectors $r$

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(Q v)_{i}=\sum_{j} p_{i j} v_{j}, \quad(r Q)_{i}=\sum_{i} p_{i j} r_{i}
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## Proposition (marginal laws)

$$
P\left(X_{n}=\cdot\right)=\mu^{0}\left(Q^{n}\right), \quad \text { for any } n \geq 0 .
$$

## A geometric interpretation of $Q^{n}$

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By definition of power of a matrix

$$
\left(Q^{n}\right)_{i j}=\sum_{k} p_{i k}\left(Q^{n-1}\right)_{k j}=\sum_{k_{1}, k_{2}, \ldots k_{(n-1)}} p_{i k_{1}} p_{k_{1} k_{2}} \cdot p_{k_{(n-1)} j}
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$$

1 Each choice of $k_{1}, \ldots, k_{(n-1)}$ defines the path of "length" $n$

$$
i \rightarrow k_{1} \rightarrow k_{2} \rightarrow \ldots \rightarrow k_{(n-1)} \rightarrow j .
$$

2. Each path has total weight given by the product of the weights.

B We sum the weights over all the paths of length $n$ joining $i$ to $j$.

$$
\begin{gathered}
\mathcal{S}=(U, E, O) \\
Q=\left(\begin{array}{lll}
1-\phi & \phi & 0 \\
\lambda & 1-\lambda-\epsilon & \epsilon \\
0 & 0 & 1
\end{array}\right) \\
Q^{2}=\left(\begin{array}{l}
\lambda \phi \\
\lambda(1-\phi)+\lambda(1-\lambda-\epsilon) \\
0
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Long time behavior

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Consider the 2 -states model, with $\lambda=0.1$ and $\phi=0.3$.


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## Problem

What can we say about $X_{n}$, for large $n$ ?
$\mathbb{P}\left(X_{n}=E\right)$ will be close to 1 . How to compute it?

$$
Q=\left(\begin{array}{cc}
0.9 & 0.1 \\
0.7 & 0.3
\end{array}\right) \quad Q^{2}=\left(\begin{array}{ll}
0.88 & 0.12 \\
0.84 & 0.16
\end{array}\right) \quad Q^{4} \sim\left(\begin{array}{cc}
0.87 & 0.13 \\
0.87 & 0.13
\end{array}\right)
$$

## Invariant distributions

Do the laws of $\left(X_{n}\right)_{n \geq 1}$ converges as $n \rightarrow \infty$ towards some limit law $\bar{\mu}$ on $\mathcal{S}$ ?

## Problem

What are relevant properties of $\bar{\mu}$ ?

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What are relevant properties of $\bar{\mu}$ ?
Since the law of $X_{n}$ is $\mu^{0}\left(Q^{n}\right)$, we should find

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\bar{\mu}=\lim _{n \rightarrow \infty} \mu^{0}\left(Q^{n}\right)
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but also

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In linear algebra notation, $\bar{\mu}$ is a row eigenvector of $Q$ with eigenvalue 1.

## Existence of Invariant distributions

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Theorem (Perron-Frobenius)
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## Proof.

Choose any $\mu^{0}$ and consider the averages

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One can prove that some limit point exists and it is invariant.
Notice that the proof gives a "way" to find $\bar{\mu}$, essentially by taking powers $Q^{n}$ and averaging.

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■ We have $P\left(X_{n+1}=i\right)=0$, for every $i \leq n$

- Any invariant probability satisifes $\bar{\mu}_{i}=0$ for every $i$


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In general the answer is NO. Let us consider one example.


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Q=\left(\begin{array}{ll}
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Question: what are the invariant distributions for $Q$ ?

## Communicating states

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Given states $i, j \in \mathcal{S}$, we say that $i \rightsquigarrow j$ if there is $n \in \mathbb{N}$ such that

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Notice that $1 \rightsquigarrow 2$ but not $2 \rightsquigarrow 1$.

## Irreducible chains

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Irreducible transition probability
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- in terms of the transition matrix, this means that for every $i, j$ there exists $m \in \mathbb{N}$ such that $\left(Q^{m}\right)_{i j}>0$.
- The name "irreducible" comes from the fact that any Markov chain can be decomposed into smaller chains which are irreducible, plus some "remainder" (called transitory states).

Regularity

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For our problem of convergence towards invariant distributions, we need something more than irreducible chains.
The problem with

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We say that the transition matrix $Q$ is regular if there exists $m \in \mathbb{N}$ such that $Q^{m}$ has strictly positive entries, i.e.

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is not regular.

In terms of the Markov chain, regularity means that there exists $m \in \mathbb{N}$ such that

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Notice that if $Q^{m}$ has strictly positive entries, then $Q^{m+1}$ also is strictly positive.

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\begin{aligned}
P\left(X_{m+1}=j \mid X_{0}=i\right) & =\sum_{k} P\left(X_{m+1}=j \mid X_{1}=k, X_{0}=i\right) P\left(X_{1}=k \mid X_{0}=i\right) \\
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because

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\sum_{k} P\left(X_{1}=k \mid X_{0}=i\right)=1 .
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## The importance of being regular

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## Theorem (Markov)

If $Q$ is a regular transition probability (on a finite state space $\mathcal{S}$ ) then there is a unique invariant distribution $\bar{\mu}$ and

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In terms of Markov chain, this answer positively to the problem

$$
P\left(X_{n}=j\right) \rightarrow \bar{\mu}(j), \quad \text { as } n \rightarrow \infty,
$$

for every $j \in \mathcal{S}$, whatever the initial law of $X_{0}$ was.

Ergodicity

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Answer to this question uses the "ergodicity" principle from physics:
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Beware: this holds for systems at equilibrium!

## Ergodicity



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In our model, the large number of individuals is encoded in the law of $X_{n}$, hence the space average at equilibrium is just

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We compute the time average of being unemployed:

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Then the ergodic principle reads

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\lim _{n \rightarrow \infty} \frac{\sharp\left\{k \in\{0,1, \ldots, n-1\} \mid X_{k}=E\right\}}{n}=\bar{\mu}(E) .
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## Ergodic theorem

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Of course, this theorem applies also to regular Markov chains.

## Continuous time Markov chains

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We can split the time $\{0,1,2, \ldots\}$ into

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\{0,1 / 4,2 / 4,3 / 4,1,5 / 4,6 / 4,7 / 4, \ldots\}
$$

and build a process $Y_{0}, Y_{1 / 4}, Y_{2 / 4}, Y_{3 / 4}, \ldots$.
If we want to have a Markov chain "close" to the original one, we should have $Q_{X} \sim Q_{Y}^{4}$, i.e. $Q_{Y}=Q_{X}^{1 / 4}$.

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If instead of 4 weeks we split into days, we should have the transition probability $Q_{X}^{1 / 30}$.

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What happens if we split times into infinitely many small intervals, i.e. continous times?
We have $Q^{1 / \infty}=Q^{0}=I d$, but imagine, for large $n$,

$$
Q^{1 / n} \sim I d+\frac{1}{n} R+\text { smaller terms }
$$

We describe the chain by means of the matrix $R$ (called transition rate matrix).

Recalling the formal expansion

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Assuming $Q^{1 / n}$ to be a stochastic matrix, then we must have

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R_{i j} \geq 0, \text { for } i \neq j, \quad R_{i j} \leq 0
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We interpret $R_{i j}$ as the rate at which we jump from $i$ to $j$.

Actually, we can describe the continuous Markov chain as follows:
1 When the particle is on the state $i \in \mathcal{S}$, take independent "alarm clocks" $A_{j}$, one for every $j \in \mathcal{S}$ with exponential laws

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An important example of such continuous time jump process has state space

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\mathcal{S}=\mathbb{N}=\{0,1, \ldots\}
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and transition rates

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Two realizations of a Poisson process:



We denote $\left(N_{t}\right)_{t \in[0, \infty)}$ the Poisson process ( $N=$ number of jumps).

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It is possible to prove that
■ $N$ is a Markov process (obvious)
■ For every $t \in[0, \infty), N_{t}$ has Poisson law

$$
P\left(N_{t}=k\right)=\frac{e^{-k}}{k!}, \quad k \in \mathbb{N}
$$

In particular,

$$
\mathbb{E}\left[N_{t}\right]=t
$$

■ the increments are independent, e.g. for $s<t<u<v$,
$N_{t}-N_{s}$ and $N_{v}-N_{u}$ are independent random variables.

One can also change the intensity of jumps (not the size), by changing the transition rates: for $\lambda>0$,

$$
R_{i j}^{\lambda}=\lambda R_{i j}
$$



