Stochastic Processes and Stochastic Calculus - 3 Markov Chains

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Overview

1 Discrete time processes

- An example
- The Markov property
- Discrete Markov chains
- Invariant distributions
- Irreducible and regular chains
- Ergodic theorems

2 Continuous time jump processes

Continuous time Markov chains

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Poisson process

We propose a simple model for

- flows between unemployment and employment
- employment and unemployment rates at "equilibrium"

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Fix a unit of time (e.g. one month) and introduce two parameters

 $\lambda :=$ probability that a worker loses his/her job in a month

 $\phi :=$ probability that an unemployed one finds a job in a month We call $\lambda, \phi \in [0, 1]$ transition probabilities.

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For n = 0, 1, 2, ..., let X_n denote the (random) state of employment of such "typical" person.

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For n = 0, 1, 2, ..., let X_n denote the (random) state of employment of such "typical" person.

We have

$$\mathbb{P}(X_{n+1} = E | X_n = E) = 1 - \lambda \quad \mathbb{P}(X_{n+1} = U | X_n = E) = \lambda$$

$$\mathbb{P}(X_{n+1} = E | X_n = U) = \phi \qquad \mathbb{P}(X_{n+1} = U | X_n = U) = 1 - \phi.$$

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NO. For example, we do not know the initial state X_0 .

The Markov assumption

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What about the probability of finding a new job, knowing that he/she has been unemployed for the last two months?

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But what about three or more? One year? At this stage we introduce the following assumption:

Markov property

At any time, regardless of the information about the past months, the next-month state of employment depends uniquely on the present one, i.e.

$$\mathbb{P}(X_{n+1} = E | X_n = U, X_{n-1} = \cdot, X_{n-2} = \cdot, \dots, X_{n-k} = \cdot) = \mathbb{P}(X_{n+1} = E | X_n = U)$$

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- The model becomes very tractable analytically
- It fits real data (?) we can make predictions
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- The model becomes very tractable analytically
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A more realistic assumption would be e.g.

$$\mathbb{P}(X_{n+1} = E | X_n = U, X_{n-1} = U, X_{n-2} = U, \dots, X_{n-k} = U) \le \mathbb{P}(X_{n+1} = E | X_n = U)$$

but how to make it quantitative?

Examples



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Examples



$$\mathbb{P}(X_2 = E, X_1 = U | X_0 = E) = \mathbb{P}(X_2 = E | X_1 = U, X_0 = E) \mathbb{P}(X_1 = U | X_0 = E)$$

= $\mathbb{P}(X_2 = E | X_1 = U) \mathbb{P}(X_1 = U | X_0 = E)$
= $\phi \lambda$

$$\mathbb{P}(X_3=E,X_2=U,X_1=U|X_0=E)=\phi(1-\phi)\lambda$$

$$\mathbb{P}(X_3 = E | X_0 = E) = ?$$

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Definition (Markov process)

Given a filtration $(\mathcal{F}_t)_{t \in T}$, a process $X_t : \Omega \to E$ is Markov if

for all $s \leq t \in T$, $A \subseteq E$ $\mathbb{P}(X_t \in A | \mathcal{F}_s) = \mathbb{P}(X_t \in A | X_s)$

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■ Usually one has $\mathcal{F}_t = \sigma(X_r : r \leq t)$, the natural filtration of X_t .

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The transition probabilities are $|S|^2$ numbers that we write matrix notation:

$$\begin{pmatrix} 1-\lambda & \lambda \\ \phi & 1-\phi \end{pmatrix} \Rightarrow Q = (p_{ij})_{i,j\in\mathcal{S}} \quad p_{ij} = \mathbb{P}(X_{n+1} = j | X_n = i) \quad \text{for } i, j \in \mathcal{S}$$

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Remark (Q is a stochastic matrix)

We must have $p_{ij} \in [0, 1]$ for $i, j \in S$ and

$$\sum_{j\in\mathcal{S}} p_{ij} = \sum_{j\in\mathcal{S}} \mathbb{P}(X_{n+1} = j | X_n = i) = \mathbb{P}(X_{n+1} \in \mathcal{S} | X_n = i) = 1$$

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■ a filtration $(\mathcal{F})_{n \in \mathbb{N}}$
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- **a** finite or countable S (set of states)
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- a filtration $(\mathcal{F})_{n \in \mathbb{N}}$

we say that a process $X_n : \Omega \to S$ ($n \in \mathbb{N}$) is a markov chain if for all $n \in \mathbb{N}$,

 $\mathbb{P}(X_{n+1} = j | \mathcal{F}_n) = \mathbb{P}(X_{n+1} = j | X_n) \quad \forall j \in S \quad (Markov property)$

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The two conditions above yield

$$\mathbb{P}(X_{n+1}=j|X_n=i,X_{n-1}=\cdot,X_{n-2}=\cdot,\ldots,X_{n-k}=\cdot)=p_{ij}$$

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Graphical representation

We associate

- **a node** to each state $i \in S$
- a weighted arrow to each transition probability p_{ij} (no arrow if $p_{ij} = 0$).

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Theorem

Given a finite or countable S, a stochastic matrix $Q = (p_{ij})_{i,j \in S}$ and a probability measure μ_0 on S.

- ∃) there exists a Markov chain $(X_n)_{n \in \mathbb{N}}$ (with respect to the natural filtration) with transition probability *Q* and law of *X*₀ equal to μ_0 .
- !) Such a chain X is unique in law, i.e. given any Markov chain Y with transition probability Q and law of X_0 equal to μ_0 , one has

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The condition X = Y in law means that for every $n \ge 1, A_0, \ldots, A_n \subseteq S$

$$\mathbb{P}(X_0 \in A_0, X_1 \in A_1, \dots, X_n \in A_n) = \mathbb{P}(Y_0 \in A_0, Y_1 \in A_1, \dots, Y_n \in A_n)$$

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How to compute the quantity

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The simplest case is that of 1-marginals, i.e. the law of X_n . For example

$$P(X_{1} = j) = \sum_{i \in S} P(X_{1} = j | X_{0} = i) P(X_{0} = i) = \sum_{i \in S} p_{ij} \mu_{i}^{0}$$

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Recall the matrix-vector products for column vectors v or row vectors r

$$(Qv)_i = \sum_j p_{ij}v_j, \quad (rQ)_i = \sum_i p_{ij}r_i$$

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If we identify μ^0 and $P(X_1 = \cdot)$ with row vectors, we have

$$P(X_1=\cdot)=\mu^0 Q$$

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Proposition (marginal laws)

$$P(X_n = \cdot) = \mu^0(Q^n)$$
, for any $n \ge 0$.

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A geometric interpretation of Q^n

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By definition of power of a matrix

$$(Q^n)_{ij} = \sum_k p_{ik} (Q^{n-1})_{kj} = \sum_{k_1, k_2, \dots, k_{(n-1)}} p_{ik_1} p_{k_1 k_2} \cdot p_{k_{(n-1)} j}$$



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1 Each choice of $k_1, \ldots, k_{(n-1)}$ defines the path of 'length' *n*

$$i \to k_1 \to k_2 \to \ldots \to k_{(n-1)} \to j.$$

2 Each path has total weight given by the product of the weights.
3 We sum the weights over all the paths of length *n* joining *i* to *j*.



Consider the 2-states model, with $\lambda = 0.1$ and $\phi = 0.3$.



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What can we say about X_n , for large n?

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 $\mathbb{P}(X_n = E)$ will be close to 1. How to compute it?

$$Q = \left(\begin{array}{cc} 0.9 & 0.1 \\ 0.7 & 0.3 \end{array}\right) \quad Q^2 = \left(\begin{array}{cc} 0.88 & 0.12 \\ 0.84 & 0.16 \end{array}\right) \quad Q^4 \sim \left(\begin{array}{cc} 0.87 & 0.13 \\ 0.87 & 0.13 \end{array}\right)$$

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Since the law of X_n is $\mu^0(Q^n)$, we should find

$$\bar{\mu} = \lim_{n \to \infty} \mu^0(Q^n)$$

but also

$$\bar{\mu}Q = \lim_{n \to \infty} \mu^0(Q^n)Q = \lim_{n \to \infty} \mu^0(Q^{n+1}) = \bar{\mu}$$

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In linear algebra notation, $\bar{\mu}$ is a row eigenvector of Q with eigenvalue 1.

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If the number of states S is finite, there exists at least one invariant distribution:

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Proof.

Choose any μ^0 and consider the averages

$$ar{\mu}^n := rac{1}{n} \left(\mu^0 + \mu^0 Q + \mu^0 Q^2 + \dots + \mu^0 Q^{n-1}
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$$\bar{\mu}^n - \bar{\mu}^n Q = \frac{1}{n} \left(\mu^0 - Q^n \mu^0 \right) \rightarrow 0$$

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Notice that the proof gives a "way" to find $\bar{\mu}$, essentially by taking powers Q^n and averaging.

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Beware: the existence result could be false if the number of states is infinite!

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• We have $P(X_{n+1} = i) = 0$, for every $i \le n$

Any invariant probability satisifes $\bar{\mu}_i = 0$ for every *i*

If the number of states is finite, we know that some $\bar{\mu}$ exists.
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Is it true that, for $j \in S$,

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In general the answer is NO. Let us consider one example.



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Question: what are the invariant distributions for Q?

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Definition (Communicating states)

Given states $i, j \in S$, we say that $i \rightsquigarrow j$ if there is $n \in \mathbb{N}$ such that

$$(Q^n)_{ij} = P(X_n = j | X_0 = i) > 0.$$

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Notice that $1 \rightsquigarrow 2$ but not $2 \rightsquigarrow 1$.

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We say that the transition matrix Q is irreducible if

for every $i, j \in S, i \rightsquigarrow j$.

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in terms of the transition matrix, this means that for every *i*, *j* there exists $m \in \mathbb{N}$ such that $(Q^m)_{ij} > 0$.

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- in terms of the transition matrix, this means that for every *i*, *j* there exists $m \in \mathbb{N}$ such that $(Q^m)_{ij} > 0$.
- The name "irreducible" comes from the fact that any Markov chain can be decomposed into smaller chains which are irreducible, plus some "remainder" (called transitory states).

For our problem of convergence towards invariant distributions, we need something more than irreducible chains. The problem with

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Regular transition probability

We say that the transition matrix Q is regular if there exists $m \in \mathbb{N}$ such that Q^m has strictly positive entries, i.e.

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This means that after some time m we are "very uncertain" about the actual position, we cannot exclude of being at any state.

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$$P(X_{m+1} = j | X_0 = i) = \sum_{k} P(X_{m+1} = j | X_1 = k, X_0 = i) P(X_1 = k | X_0 = i)$$
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because

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Theorem (Markov)

If Q is a regular transition probability (on a finite state space S) then there is a unique invariant distribution $\bar{\mu}$ and

 $\lim_{n\to\infty}(\boldsymbol{Q}^n)_{ij}=\bar{\mu}(j),\quad\text{for every }i,j\in\mathcal{S}.$

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In terms of Markov chain, this answer positively to the problem

$$P(X_n = j) \rightarrow \overline{\mu}(j)$$
, as $n \rightarrow \infty$,

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for every $j \in S$, whatever the initial law of X_0 was.

Going back to our model of employment flows,



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What is the "typical" time spent looking for a new job (over a long period)?

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Answer to this question uses the "ergodicity" principle from physics:

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Beware: this holds for systems at equilibrium!



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In our model, the large number of individuals is encoded in the law of X_n , hence the space average at equilibrium is just

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$$\sharp \{k \in \{0, 1, \dots, n-1\} \mid X_k = E\}$$

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$$\frac{\#\{k \in \{0, 1, \dots, n-1\} \mid X_k = E\}}{n}$$

Then the ergodic principle reads

$$\lim_{n \to \infty} \frac{\#\{k \in \{0, 1, \dots, n-1\} \mid X_k = E\}}{n} = \bar{\mu}(E).$$
The ergodic principle is actually a theorem in the setting of Markov chains.

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Ergodic theorem

Let $(X_n)_{n\geq 1}$, be a irreducible Markov chain on a finite state space S. Then

$$\lim_{n\to\infty}\frac{\sharp\left\{k\in\{0,1,\ldots,n-1\}\mid X_k=j\right\}}{n}=\bar{\mu}(j),\quad\text{for every }j\in\mathcal{S}.$$

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Of course, this theorem applies also to regular Markov chains.

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How to modify our model if, instead of measuring time intervals in months we have better precision, e.g. weeks (1 month \sim 4 weeks).

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and build a process $Y_0, Y_{1/4}, Y_{2/4}, Y_{3/4}, \ldots$

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What happens if we split times into infinitely many small intervals, i.e. continous times?

We have $Q^{1/\infty} = Q^0 = Id$, but imagine, for large *n*,

$$Q^{1/n} \sim Id + \frac{1}{n}R +$$
smaller terms.

We describe the chain by means of the matrix R (called transition rate matrix).

Recalling the formal expansion

$$Q^{1/n} \sim Id + \frac{1}{n}R + \text{smaller terms.}$$

Assuming $Q^{1/n}$ to be a stochastic matrix, then we must have

$$R_{ij} \ge 0$$
, for $i \ne j$, $R_{ii} \le 0$

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$$\sum_{j}R_{ij}=0.$$

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We interpret R_{ij} as the rate at which we jump from *i* to *j*.

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Actually, we can describe the continuous Markov chain as follows:

When the particle is on the state $i \in S$, take independent "alarm clocks" A_i , one for every $j \in S$ with exponential laws

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An important example of such continuous time jump process has state space

$$\mathcal{S} = \mathbb{N} = \{0, 1, \ldots\}$$

and transition rates

$$R_{ij} = \begin{cases} 1 & \text{if } j = i+1 \\ -1 & \text{if } j = i \\ 0 & \text{otherwise.} \end{cases}$$

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Two realizations of a Poisson process:



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We denote $(N_t)_{t \in [0,\infty)}$ the Poisson process (N = number of jumps).



We denote $(N_t)_{t \in [0,\infty)}$ the Poisson process (N = number of jumps). It is possible to prove that

- N is a Markov process (obvious)
- For every $t \in [0, \infty)$, N_t has Poisson law

$$P(N_t = k) = rac{e^{-k}}{k!}, \quad k \in \mathbb{N}.$$

In particular,

$$\mathbb{E}\left[N_t\right]=t.$$

• the increments are independent, e.g. for s < t < u < v,

 $N_t - N_s$ and $N_v - N_u$ are independent random variables.

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One can also change the intensity of jumps (not the size), by changing the transition rates: for $\lambda > 0$,

$$R_{ij}^{\lambda} = \lambda R_{ij}$$



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